

Superconcentration inequalities for centered Gaussian stationary processes

Kevin Tanguy

Université de Toulouse

January 26, 2016

- ▶ What is superconcentration ?
- ▶ Convergence of extremes (Gaussian case).
- ▶ Superconcentration inequality for stationary Gaussian sequences.
- ▶ Main result (abstract theorem)
- ▶ Tools and sketch of the proof.

- ▶ What is superconcentration ?

$$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$$

$$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma) \text{ with } \mathbb{E}[X_i^2] = 1.$$

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ with $\mathbb{E}[X_i^2] = 1$.

Set $M_n = \max_i X_i$.

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ with $\mathbb{E}[X_i^2] = 1$.

Set $M_n = \max_i X_i$.

Variance upper bound

$$\text{Var}(M_n) \leq ?$$

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ with $\mathbb{E}[X_i^2] = 1$.

Set $M_n = \max_i X_i$.

Variance upper bound

$$\text{Var}(M_n) \leq ?$$

Classical concentration theory

$$\text{Var}(M_n) \leq \max_i \text{Var}(X_i).$$

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ with $\mathbb{E}[X_i^2] = 1$.

Set $M_n = \max_i X_i$.

Variance upper bound

$$\text{Var}(M_n) \leq ?$$

Classical concentration theory

$$\text{Var}(M_n) \leq \max_i \text{Var}(X_i).$$

sharp inequality,

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ with $\mathbb{E}[X_i^2] = 1$.

Set $M_n = \max_i X_i$.

Variance upper bound

$$\text{Var}(M_n) \leq ?$$

Classical concentration theory

$$\text{Var}(M_n) \leq \max_i \text{Var}(X_i).$$

sharp inequality, does not depend on Γ .

Take $\Gamma = Id$ (X_i 's independent).

Take $\Gamma = Id$ (X_i 's independent).

$$\text{Var}(M_n) \leq 1$$

Take $\Gamma = Id$ (X_i 's independent).

$$\text{Var}(M_n) \leq 1$$

In fact,

$$\text{Var}(M_n) \leq \frac{C}{\log n}, \quad C > 0$$

Classical theory suboptimal

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Lot of different models

- ▶ Largest eigenvalue in random matrix theory.

Largest eigenvalue in random matrix theory.

- ▶ M random matrix from the GUE, namely

$$\mathbb{P}(dM) = Z_N^{-1} \exp(-\text{Tr}(M^2)/2\sigma^2) dM, \quad \sigma^2 = 1/4N$$

Largest eigenvalue in random matrix theory.

- ▶ M random matrix from the GUE, namely

$$\mathbb{P}(dM) = Z_N^{-1} \exp(-\text{Tr}(M^2)/2\sigma^2) dM, \quad \sigma^2 = 1/4N$$

- ▶ $N^{2/3}(\lambda_{\max} - 1) \xrightarrow{d} TW$, ($N \rightarrow \infty$), where TW so-called Tracy-Widom distribution. [Tracy-Widom '90]

Largest eigenvalue in random matrix theory.

- ▶ M random matrix from the GUE, namely

$$\mathbb{P}(dM) = Z_N^{-1} \exp(-\text{Tr}(M^2)/2\sigma^2) dM, \quad \sigma^2 = 1/4N$$

- ▶ $N^{2/3}(\lambda_{\max} - 1) \xrightarrow{d} TW$, ($N \rightarrow \infty$), where TW so-called Tracy-Widom distribution. [Tracy-Widom '90]

- ▶ Classical theory : $\text{Var}(\lambda_{\max}) \leq \frac{C}{N}$

Largest eigenvalue in random matrix theory.

- ▶ M random matrix from the GUE, namely

$$\mathbb{P}(dM) = Z_N^{-1} \exp(-\text{Tr}(M^2)/2\sigma^2) dM, \quad \sigma^2 = 1/4N$$

- ▶ $N^{2/3}(\lambda_{\max} - 1) \xrightarrow{d} TW$, ($N \rightarrow \infty$), where TW so-called Tracy-Widom distribution. [Tracy-Widom '90]

- ▶ Classical theory : $\text{Var}(\lambda_{\max}) \leq \frac{C}{N}$
- ▶ In fact, $\text{Var}(\lambda_{\max}) \leq \frac{C}{N^{4/3}}$ [Ledoux-Rider '10].

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Lot of different models

- ▶ Largest eigenvalue in random matrix theory.

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Lot of different models

- ▶ Largest eigenvalue in random matrix theory.
- ▶ Branching random walk.

Branching random walk.

- ▶ Take a binary tree of depth N .

Branching random walk.

- ▶ Take a binary tree of depth N .
- ▶ Put X_e *i.i.d.* $\mathcal{N}(0, 1)$ on each edge e .

Branching random walk.

- ▶ Take a binary tree of depth N .
- ▶ Put X_e *i.i.d.* $\mathcal{N}(0, 1)$ on each edge e .
- ▶ Take a path π from the top to the bottom of the tree,
$$X_\pi = \sum_{e \in \pi} X_e.$$

Branching random walk.

- ▶ Take a binary tree of depth N .
 - ▶ Put X_e *i.i.d.* $\mathcal{N}(0, 1)$ on each edge e .
 - ▶ Take a path π from the top to the bottom of the tree,
 $X_\pi = \sum_{e \in \pi} X_e$.
- ▶ Classical theory : $\text{Var}(\max_\pi X_\pi) \leq N$

Branching random walk.

- ▶ Take a binary tree of depth N .
- ▶ Put X_e *i.i.d.* $\mathcal{N}(0, 1)$ on each edge e .
- ▶ Take a path π from the top to the bottom of the tree,
 $X_\pi = \sum_{e \in \pi} X_e$.

- ▶ Classical theory : $\text{Var}(\max_\pi X_\pi) \leq N$ ($X_\pi \sim \mathcal{N}(0, N)$).

Branching random walk.

- ▶ Take a binary tree of depth N .
- ▶ Put X_e *i.i.d.* $\mathcal{N}(0, 1)$ on each edge e .
- ▶ Take a path π from the top to the bottom of the tree,
 $X_\pi = \sum_{e \in \pi} X_e$.

- ▶ Classical theory : $\text{Var}(\max_\pi X_\pi) \leq N$ ($X_\pi \sim \mathcal{N}(0, N)$).
- ▶ **In fact, $\text{Var}(\max_\pi X_\pi) \leq C$** [Bramson-Ding-Zeitouni].

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Lot of different models

- ▶ Largest eigenvalue in random matrix theory.
- ▶ Branching random walk.

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Lot of different models

- ▶ Largest eigenvalue in random matrix theory.
- ▶ Branching random walk.
- ▶ Discrete Gaussian free field on \mathbb{Z}^d .

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Lot of different models

- ▶ Largest eigenvalue in random matrix theory.
- ▶ Branching random walk.
- ▶ Discrete Gaussian free field on \mathbb{Z}^d .
- ▶ Free energy in spin glasses theory (SK model).

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Lot of different models

- ▶ Largest eigenvalue in random matrix theory.
- ▶ Branching random walk.
- ▶ Discrete Gaussian free field on \mathbb{Z}^d .
- ▶ Free energy in spin glasses theory (SK model).
- ▶ First passage in percolation theory.

Classical theory suboptimal

Chatterjee's terminology : superconcentration phenomenon.

Lot of different models

- ▶ Largest eigenvalue in random matrix theory.
- ▶ Branching random walk.
- ▶ Discrete Gaussian free field on \mathbb{Z}^d .
- ▶ Free energy in spin glasses theory (SK model).
- ▶ First passage in percolation theory.
- ▶ Stationary Gaussian sequences.

- ▶ What is superconcentration ?
- ▶ Convergence of extremes (Gaussian case).

Stationary Gaussian sequences

Theorem [Berman '64]

$(X_i)_{i \geq 0}$ centered normalized stationary Gaussian sequence with covariance function ϕ .

Stationary Gaussian sequences

Theorem [Berman '64]

$(X_i)_{i \geq 0}$ centered normalized stationary Gaussian sequence with covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$),

Stationary Gaussian sequences

Theorem [Berman '64]

$(X_i)_{i \geq 0}$ centered normalized stationary Gaussian sequence with covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$), then

$$a_n(M_n - b_n) \xrightarrow{d} G, \quad (n \rightarrow \infty)$$

Stationary Gaussian sequences

Theorem [Berman '64]

$(X_i)_{i \geq 0}$ centered normalized stationary Gaussian sequence with covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$), then

$$a_n(M_n - b_n) \xrightarrow{d} G, \quad (n \rightarrow \infty)$$

where $a_n = \sqrt{2 \log n}$

Stationary Gaussian sequences

Theorem [Berman '64]

$(X_i)_{i \geq 0}$ centered normalized stationary Gaussian sequence with covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$), then

$$a_n(M_n - b_n) \xrightarrow{d} G, \quad (n \rightarrow \infty)$$

where $a_n = \sqrt{2 \log n}$ and $\mathbb{P}(G \geq t) = 1 - e^{-e^{-t}}$, $t \in \mathbb{R}$ (Gumbel).

Stationary Gaussian sequences

Theorem [Berman '64]

$(X_i)_{i \geq 0}$ centered normalized stationary Gaussian sequence with covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$), then

$$a_n(M_n - b_n) \xrightarrow{d} G, \quad (n \rightarrow \infty)$$

where $a_n = \sqrt{2 \log n}$ and $\mathbb{P}(G \geq t) = 1 - e^{-e^{-t}}$, $t \in \mathbb{R}$ (Gumbel).

Note : $\mathbb{P}(G \geq t) \sim e^{-t}$

Results from classical theory

$$\text{Var}(M_n) \leq 1.$$

Results from classical theory

$$\text{Var}(M_n) \leq 1.$$

Question :

- ▶ Correct bound ?

Results from classical theory

$$\text{Var}(M_n) \leq 1.$$

Question :

- ▶ Correct bound ?

Proposition [Chatterjee '14]

$$\text{Var}(M_n) \leq \frac{C}{\log n}.$$

- ▶ What is superconcentration ?
- ▶ Convergence of extremes (Gaussian case).
- ▶ Superconcentration inequality for stationary Gaussian sequences.

(Super)concentration inequality ?

(Super)concentration inequality ?

Gaussian concentration inequality [Borel-Sudakov-Tsirelson '76]

Take $X \sim \mathcal{N}(0, Id)$ and $F : \mathbb{R}^n \rightarrow \mathbb{R}$ Lipschitz.

(Super)concentration inequality ?

Gaussian concentration inequality [Borel-Sudakov-Tsirelson '76]

Take $X \sim \mathcal{N}(0, Id)$ and $F : \mathbb{R}^n \rightarrow \mathbb{R}$ Lipschitz.

$$\mathbb{P}\left(|F(X) - \mathbb{E}[F(X)]| \geq t\right) \leq 2e^{-t^2/2\|F\|_{Lip}^2}, t \geq 0$$

If $F(x) = \max_i x_i$, $\|F\|_{Lip}^2 \leq 1$

(Super)concentration inequality ?

Gaussian concentration inequality [Borel-Sudakov-Tsirelson '76]

Take $X \sim \mathcal{N}(0, Id)$ and $F : \mathbb{R}^n \rightarrow \mathbb{R}$ Lipschitz.

$$\mathbb{P}\left(|F(X) - \mathbb{E}[F(X)]| \geq t\right) \leq 2e^{-t^2/2\|F\|_{Lip}^2}, t \geq 0$$

If $F(x) = \max_i x_i$, $\|F\|_{Lip}^2 \leq 1$

$$\mathbb{P}\left(a_n |M_n - \mathbb{E}[M_n]| \geq t\right) \leq 2e^{-\frac{t^2}{2a_n^2}}, t \geq 0, a_n = \sqrt{2 \log n}$$

(Super)concentration inequality ?

Gaussian concentration inequality [Borel-Sudakov-Tsirelson '76]

Take $X \sim \mathcal{N}(0, Id)$ and $F : \mathbb{R}^n \rightarrow \mathbb{R}$ Lipschitz.

$$\mathbb{P}\left(|F(X) - \mathbb{E}[F(X)]| \geq t\right) \leq 2e^{-t^2/2\|F\|_{Lip}^2}, t \geq 0$$

If $F(x) = \max_i x_i$, $\|F\|_{Lip}^2 \leq 1$

$$\mathbb{P}\left(a_n |M_n - \mathbb{E}[M_n]| \geq t\right) \leq 2e^{-\frac{t^2}{2a_n^2}}, t \geq 0, a_n = \sqrt{2 \log n}$$

does not reflect Gumbel asymptotics.

Question : (super)concentration inequality reflecting convergence of extremes ?

Question : **(super)concentration inequality reflecting convergence of extremes ?**

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence,

Question : **(super)concentration inequality reflecting convergence of extremes ?**

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence, covariance function ϕ .

Question : **(super)concentration inequality reflecting convergence of extremes ?**

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence, covariance function ϕ .
Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and

Question : **(super)concentration inequality reflecting convergence of extremes ?**

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence, covariance function ϕ .
Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and technicals hypothesis, then

Question : **(super)concentration inequality reflecting convergence of extremes ?**

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence, covariance function ϕ .
Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and technicals hypothesis, then

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct\sqrt{\log n}}, \quad t \geq 0.$$

Question : **(super)concentration inequality reflecting convergence of extremes ?**

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence, covariance function ϕ .
Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and technical hypothesis, then

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct\sqrt{\log n}}, \quad t \geq 0.$$

Remark : same inequality holds with b_n instead of $\mathbb{E}[M_n]$.

Question : **(super)concentration inequality reflecting convergence of extremes ?**

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence, covariance function ϕ .
Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and technicals hypothesis, then

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct\sqrt{\log n}}, \quad t \geq 0.$$

Remark : same inequality holds with b_n instead of $\mathbb{E}[M_n]$.

$$\mathbb{P}\left(\sqrt{\log n}|M_n - b_n| \geq t\right) \leq Ce^{-ct}, \quad t \geq 0.$$

Question : **(super)concentration inequality reflecting convergence of extremes ?**

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence, covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and technicals hypothesis, then

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct\sqrt{\log n}}, \quad t \geq 0.$$

Remark : same inequality holds with b_n instead of $\mathbb{E}[M_n]$.

$$\mathbb{P}\left(\sqrt{\log n}|M_n - b_n| \geq t\right) \leq Ce^{-ct}, \quad t \geq 0.$$

► Reflects asymptotics Gumbel

Recall

Theorem [Berman 64']

$(X_i)_{i \geq 0}$ centered normalized stationary Gaussian sequence with covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$), then

$$a_n(M_n - b_n) \xrightarrow{d} G, \quad (n \rightarrow \infty)$$

where $a_n = \sqrt{2 \log n}$ and $\mathbb{P}(G \geq t) = 1 - e^{-e^{-t}}$, $t \in \mathbb{R}$ (Gumbel).

Note : $\mathbb{P}(G \geq t) \sim e^{-t}$

Theorem [T. 15']

$(X_i)_{i \geq 0}$ centered stationary Gaussian, covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and technical hypothesis, then

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct\sqrt{\log n}}, \quad t \geq 0.$$

Remark : same inequality holds with b_n instead of $\mathbb{E}[M_n]$.

- ▶ Reflects Gumbel asymptotics.

Theorem [T. 15']

$(X_i)_{i \geq 0}$ centered stationary Gaussian, covariance function ϕ . Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and technical hypothesis, then

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct\sqrt{\log n}}, \quad t \geq 0.$$

Remark : same inequality holds with b_n instead of $\mathbb{E}[M_n]$.

- ▶ Reflects Gumbel asymptotics.
- ▶ Implies $\text{Var}(\max_i X_i) \leq \frac{C}{\log n}$ (optimal).

Proof ?

Proof ?

Chatterjee's scheme of proof for the variance at the exponential level.

Proof ?

Chatterjee's scheme of proof for the variance at the exponential level.

General theorem implies superconcentration inequality for Gaussian stationary sequences.

- ▶ What is superconcentration ?
- ▶ Convergence of extremes (Gaussian case).
- ▶ Superconcentration inequality for stationary Gaussian sequences.
- ▶ Main result (abstract theorem)

General theorem [T. '15]

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

General theorem [T. '15]

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;

General theorem [T. '15]

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;

Explanation : if $\Gamma = Id$, choose $r_0 > 0$ then $\mathcal{C}(r_0) = \{\{1\}, \dots, \{n\}\}$.

General theorem [T. '15]

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;

Explanation : if $\Gamma = Id$, choose $r_0 > 0$ then $\mathcal{C}(r_0) = \{\{1\}, \dots, \{n\}\}$.
Indeed, if $\Gamma_{ij} > 0$ then $i = j$.

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, a.s., $\sum_{D \in \mathcal{C}(r_0)} \mathbf{1}_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, *a.s.*, $\sum_{D \in \mathcal{C}(r_0)} 1_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

Explanation : $\mathcal{C}(r_0) = \{\{1\}, \dots, \{n\}\}$ partition of $\{1, \dots, n\}$, so $\sum_i 1_{\{I=i\}} = 1$.

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, a.s., $\sum_{D \in \mathcal{C}(r_0)} 1_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

Explanation : $\mathcal{C}(r_0) = \left\{ \{1\}, \dots, \{n\} \right\}$ partition of $\{1, \dots, n\}$, so

$\sum_i 1_{\{I=i\}} = 1$. In general, $\mathcal{C}(r_0)$ "slightly bigger" than a partition and $C \geq 1$.

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, a.s., $\sum_{D \in \mathcal{C}(r_0)} \mathbf{1}_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

Let $\rho(r_0) = \max_{D \in \mathcal{C}(r_0)} \mathbb{P}(I \in D)$.

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, a.s., $\sum_{D \in \mathcal{C}(r_0)} \mathbf{1}_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

Let $\rho(r_0) = \max_{D \in \mathcal{C}(r_0)} \mathbb{P}(I \in D)$.

Explanation : If $\Gamma = Id$, $\mathcal{C}(r_0) = \{\{1\}, \dots, \{n\}\}$,

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, a.s., $\sum_{D \in \mathcal{C}(r_0)} \mathbf{1}_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

Let $\rho(r_0) = \max_{D \in \mathcal{C}(r_0)} \mathbb{P}(I \in D)$.

Explanation : If $\Gamma = Id$, $\mathcal{C}(r_0) = \{\{1\}, \dots, \{n\}\}$,
 $\mathbb{P}(I = i) = \mathbb{P}(X_i \geq X_j \forall j) =$

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, a.s., $\sum_{D \in \mathcal{C}(r_0)} \mathbf{1}_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

Let $\rho(r_0) = \max_{D \in \mathcal{C}(r_0)} \mathbb{P}(I \in D)$.

Explanation : If $\Gamma = Id$, $\mathcal{C}(r_0) = \{\{1\}, \dots, \{n\}\}$,
 $\mathbb{P}(I = i) = \mathbb{P}(X_i \geq X_j \forall j) = 1/n$.

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, a.s., $\sum_{D \in \mathcal{C}(r_0)} \mathbf{1}_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

Let $\rho(r_0) = \max_{D \in \mathcal{C}(r_0)} \mathbb{P}(I \in D)$.

Explanation : If $\Gamma = Id$, $\mathcal{C}(r_0) = \{\{1\}, \dots, \{n\}\}$,

$\mathbb{P}(I = i) = \mathbb{P}(X_i \geq X_j \forall j) = 1/n$. Then $\rho(r_0) \leq 1/n$.

General theorem [T. 15']

$X = (X_1, \dots, X_n) \sim \mathcal{N}(0, \Gamma)$ Assume that for some $r_0 \geq 0$, there exists a covering $\mathcal{C}(r_0)$ of $\{1, \dots, n\}$ verifying :

- ▶ for all $i, j \in \{1, \dots, n\}$ such that $\Gamma_{ij} \geq r_0$, there exists $D \in \mathcal{C}(r_0)$ such that $i, j \in D$;
- ▶ there exists $C \geq 1$ such that, *a.s.*, $\sum_{D \in \mathcal{C}(r_0)} \mathbf{1}_{\{I \in D\}} \leq C$, where $I = \operatorname{argmax}_i X_i$.

Let $\rho(r_0) = \max_{D \in \mathcal{C}(r_0)} \mathbb{P}(I \in D)$. Then, for every $\theta \in \mathbb{R}$,

$$\operatorname{Var}\left(e^{\theta M_n/2}\right) \leq C \frac{\theta^2}{4} \left(r_0 + \frac{1}{\log(1/\rho(r_0))} \right) \mathbb{E}\left[e^{\theta M_n}\right].$$

Conclusion ?

Exponential Poincaré inequality

Conclusion ?

Take Z random variable

Exponential Poincaré inequality

Conclusion ?

Take Z random variable

$$\text{Var} \left(e^{\theta Z/2} \right) \leq \frac{\theta^2}{4} \mathbf{K}\mathbf{E} \left[e^{\theta Z} \right], \quad \theta \in \mathbb{R}$$

Exponential Poincaré inequality

Conclusion ?

Take Z random variable

$$\text{Var} \left(e^{\theta Z/2} \right) \leq \frac{\theta^2}{4} \mathbb{K} \mathbb{E} \left[e^{\theta Z} \right], \quad \theta \in \mathbb{R}$$

implies

$$\mathbb{P} \left(|Z - \mathbb{E}[Z]| \geq t \right) \leq 6e^{-ct/\sqrt{K}}, \quad t > 0$$

So

$$\text{Var}\left(e^{\theta M_n/2}\right) \leq C \frac{\theta^2}{4} \left(r_0 + \frac{1}{\log(1/\rho(r_0))} \right) \mathbb{E}\left[e^{\theta M_n}\right].$$

So

$$\text{Var}\left(e^{\theta M_n/2}\right) \leq C \frac{\theta^2}{4} \left(r_0 + \frac{1}{\log(1/\rho(r_0))}\right) \mathbb{E}\left[e^{\theta M_n}\right].$$

implies

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct/\sqrt{K_{r_0}}}, \quad t \geq 0,$$

So

$$\text{Var}\left(e^{\theta M_n/2}\right) \leq C \frac{\theta^2}{4} \left(r_0 + \frac{1}{\log(1/\rho(r_0))}\right) \mathbb{E}\left[e^{\theta M_n}\right].$$

implies

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct/\sqrt{K_{r_0}}}, \quad t \geq 0,$$

where $K_{r_0} = \max\left(r_0, \frac{1}{\log(1/\rho(r_0))}\right)$ and $c > 0$.

So

$$\text{Var}\left(e^{\theta M_n/2}\right) \leq C \frac{\theta^2}{4} \left(r_0 + \frac{1}{\log(1/\rho(r_0))} \right) \mathbb{E}\left[e^{\theta M_n}\right].$$

implies

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct/\sqrt{K_{r_0}}}, \quad t \geq 0,$$

where $K_{r_0} = \max\left(r_0, \frac{1}{\log(1/\rho(r_0))}\right)$ and $c > 0$.

(Stationary case : $K_{r_0} = 1/\log n$)

Recall

Theorem [T. '15]

$(X_i)_{i \geq 0}$ centered stationary Gaussian sequence, covariance function ϕ .
Assume $\phi(n) = o(\log n)$ ($n \rightarrow \infty$) and technical hypothesis, then

$$\mathbb{P}(|M_n - \mathbb{E}[M_n]| \geq t) \leq 6e^{-ct\sqrt{\log n}}, \quad t \geq 0.$$

- ▶ What is superconcentration ?
- ▶ Convergence of extremes (Gaussian case).
- ▶ Superconcentration inequality for stationary Gaussian sequences.
- ▶ Main result (abstract theorem)
- ▶ Tools and sketch of the proof.

Main steps

- ▶ Semigroup representation of the variance

Main steps

- ▶ Semigroup representation of the variance
- ▶ Hypercontractivity

Main steps

- ▶ Semigroup representation of the variance
- ▶ Hypercontractivity
- ▶ Proper use of the covering $\mathcal{C}(r_0)$.

$$X \sim \mathcal{N}(0, \Gamma)$$

$X \sim \mathcal{N}(0, \Gamma)$, take Y independent copy of X

$X \sim \mathcal{N}(0, \Gamma)$, take Y independent copy of X

Ornstein Uhlenbeck generalized

$$X^t = Xe^{-t} + \sqrt{1 - e^{-2t}}Y, \quad t \geq 0$$

$X \sim \mathcal{N}(0, \Gamma)$, take Y independent copy of X

Ornstein Uhlenbeck generalized

$$X^t = Xe^{-t} + \sqrt{1 - e^{-2t}}Y, \quad t \geq 0$$

$$Q_t f(x) = \mathbb{E} \left[f \left(xe^{-t} + \sqrt{1 - e^{-2t}}Y \right) \right], \quad t \geq 0, x \in \mathbb{R}^n$$

$X \sim \mathcal{N}(0, \Gamma)$, take Y independent copy of X

Ornstein Uhlenbeck generalized

$$X^t = Xe^{-t} + \sqrt{1 - e^{-2t}}Y, \quad t \geq 0$$

$$Q_t f(x) = \mathbb{E} \left[f \left(xe^{-t} + \sqrt{1 - e^{-2t}}Y \right) \right], \quad t \geq 0, x \in \mathbb{R}^n$$

$(Q_t)_{t \geq 0}$ is hypercontractive :

$X \sim \mathcal{N}(0, \Gamma)$, take Y independent copy of X

Ornstein Uhlenbeck generalized

$$X^t = Xe^{-t} + \sqrt{1 - e^{-2t}}Y, \quad t \geq 0$$

$$Q_t f(x) = \mathbb{E} \left[f \left(xe^{-t} + \sqrt{1 - e^{-2t}}Y \right) \right], \quad t \geq 0, x \in \mathbb{R}^n$$

$(Q_t)_{t \geq 0}$ is hypercontractive :

$$\mathbb{E} \left[|Q_t f|^2 \right]^{1/2} \leq \mathbb{E} [|f|^p]^{1/p}, \quad p = 1 + e^{-2t} < 2.$$

Variance semigroup representation

$$X \sim \mathcal{N}(0, \Gamma)$$

Variance semigroup representation

$X \sim \mathcal{N}(0, \Gamma)$

When $\Gamma = Id$,

$$\text{Var}(f) = \int_0^\infty e^{-t} \mathbb{E} [\nabla f \cdot Q_t \nabla f] dt$$

Variance semigroup representation

$X \sim \mathcal{N}(0, \Gamma)$

When $\Gamma = Id$,

$$\text{Var}(f) = \int_0^\infty e^{-t} \mathbb{E} [\nabla f \cdot Q_t \nabla f] dt$$

General case

$$\text{Var}(e^{\theta M_n/2}) = \int_0^\infty e^{-t} \sum_{i,j} \Gamma_{ij} \mathbb{E} [f_i Q_t f_j] dt$$

Variance semigroup representation

$X \sim \mathcal{N}(0, \Gamma)$

When $\Gamma = Id$,

$$\text{Var}(f) = \int_0^\infty e^{-t} \mathbb{E} [\nabla f \cdot Q_t \nabla f] dt$$

General case

$$\text{Var}(e^{\theta M_n/2}) = \int_0^\infty e^{-t} \sum_{i,j} \Gamma_{ij} \mathbb{E} [f_i Q_t f_j] dt$$

with $f_i = \partial_i (e^{\theta M_n/2})$

Variance semigroup representation

$$X \sim \mathcal{N}(0, \Gamma)$$

When $\Gamma = Id$,

$$\text{Var}(f) = \int_0^\infty e^{-t} \mathbb{E} [\nabla f \cdot Q_t \nabla f] dt$$

General case

$$\text{Var}(e^{\theta M_n/2}) = \int_0^\infty e^{-t} \sum_{i,j} \Gamma_{ij} \mathbb{E} [f_i Q_t f_j] dt$$

with $f_i = \partial_i (e^{\theta M_n/2}) = \frac{\theta}{2} \mathbf{1}_{\{X_i = \max\}} e^{\theta M_n/2}$.

Variance semigroup representation

$$X \sim \mathcal{N}(0, \Gamma)$$

When $\Gamma = Id$,

$$\text{Var}(f) = \int_0^\infty e^{-t} \mathbb{E}[\nabla f \cdot Q_t \nabla f] dt$$

General case

$$\text{Var}(e^{\theta M_n/2}) = \int_0^\infty e^{-t} \sum_{i,j} \Gamma_{ij} \mathbb{E}[f_i Q_t f_j] dt$$

with $f_i = \partial_i(e^{\theta M_n/2}) = \frac{\theta}{2} \mathbf{1}_{\{X_i = \max\}} e^{\theta M_n/2}$.

Similar as

$$\partial_i \left(\max_j x_j \right) = \partial_i \left(\sum_j x_j \mathbf{1}_{\{x_j = \max\}} \right)$$

Variance semigroup representation

$$X \sim \mathcal{N}(0, \Gamma)$$

When $\Gamma = Id$,

$$\text{Var}(f) = \int_0^\infty e^{-t} \mathbb{E} [\nabla f \cdot Q_t \nabla f] dt$$

General case

$$\text{Var}(e^{\theta M_n/2}) = \int_0^\infty e^{-t} \sum_{i,j} \Gamma_{ij} \mathbb{E} [f_i Q_t f_j] dt$$

with $f_i = \partial_i (e^{\theta M_n/2}) = \frac{\theta}{2} \mathbf{1}_{\{X_i = \max\}} e^{\theta M_n/2}$.

Similar as

$$\partial_i \left(\max_j x_j \right) = \partial_i \left(\sum_j x_j \mathbf{1}_{\{x_j = \max\}} \right) = \mathbf{1}_{\{x_i = \max\}}$$

Key steps for the proof

Denote $l = \operatorname{argmax}_i X_i$ and $l^t = \operatorname{argmax}_i X_i^t$.

Key steps for the proof

Denote $l = \operatorname{argmax}_i X_i$ and $l^t = \operatorname{argmax}_i X_i^t$.

$$\mathcal{I} = \mathbb{E} \left[\sum_{i,j=1}^n \Gamma_{ij} \mathbf{1}_{\{X_i = \max\}} e^{\theta M_n/2} \mathbf{1}_{\{X_j^t = \max\}} e^{\theta M_n^t/2} \right]$$

Key steps for the proof

Denote $I = \operatorname{argmax}_i X_i$ and $I^t = \operatorname{argmax}_i X_i^t$.

$$\begin{aligned}\mathcal{I} &= \mathbb{E} \left[\sum_{i,j=1}^n \Gamma_{ij} \mathbf{1}_{\{X_i = \max\}} e^{\theta M_n/2} \mathbf{1}_{\{X_j^t = \max\}} e^{\theta M_n^t/2} \right] \\ &= \mathbb{E} \left[e^{\theta(M_n + M_n^t)/2} \sum_{i,j=1}^n \Gamma_{ij} \mathbf{1}_{\{I=i\}} \mathbf{1}_{\{I^t=j\}} \right]\end{aligned}$$

Key steps for the proof

Denote $I = \operatorname{argmax}_i X_i$ and $I^t = \operatorname{argmax}_i X_i^t$.

$$\begin{aligned}\mathcal{I} &= \mathbb{E} \left[\sum_{i,j=1}^n \Gamma_{ij} \mathbf{1}_{\{X_i = \max\}} e^{\theta M_n/2} \mathbf{1}_{\{X_j^t = \max\}} e^{\theta M_n^t/2} \right] \\ &= \mathbb{E} \left[e^{\theta(M_n + M_n^t)/2} \sum_{i,j=1}^n \Gamma_{ij} \mathbf{1}_{\{I=i\}} \mathbf{1}_{\{I^t=j\}} \right] \\ &= \mathbb{E} \left[e^{\theta(M_n + M_n^t)/2} \Gamma_{II^t} \right]\end{aligned}$$

Key steps for the proof

Denote $l = \operatorname{argmax}_i X_i$ and $l^t = \operatorname{argmax}_i X_i^t$.

$$\begin{aligned}\mathcal{I} &= \mathbb{E} \left[\sum_{i,j=1}^n \Gamma_{ij} \mathbf{1}_{\{X_i = \max\}} e^{\theta M_n/2} \mathbf{1}_{\{X_j^t = \max\}} e^{\theta M_n^t/2} \right] \\ &= \mathbb{E} \left[e^{\theta(M_n + M_n^t)/2} \sum_{i,j=1}^n \Gamma_{ij} \mathbf{1}_{\{l=i\}} \mathbf{1}_{\{l^t=j\}} \right] \\ &= \mathbb{E} \left[e^{\theta(M_n + M_n^t)/2} \Gamma_{ll^t} \right] \\ &= \sum_{k \geq 0} \mathbb{E} \left[e^{\theta(M_n + M_n^t)/2} \Gamma_{ll^t} \mathbf{1}_{\{2^{-k-1} \leq \Gamma_{ll^t} \leq 2^{-k}\}} \right].\end{aligned}$$

Key steps for the proof

Cut the sum according to the size of Γ_{ij} .

Key steps for the proof

Cut the sum according to the size of Γ_{ij} .

Set $F^t = M_n + M_n^t$ and $k_0 = \min\{k \geq 0; r_0 \leq 2^{-k-1}\}$.

Key steps for the proof

Cut the sum according to the size of Γ_{ij} .

Set $F^t = M_n + M_n^t$ and $k_0 = \min\{k \geq 0; r_0 \leq 2^{-k-1}\}$.

$$\mathcal{I} \leq \sum_{k=0}^{k_0} 2^{-k} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{\Gamma \geq r_0\}} \right] + \sum_{k=k_0+1}^{\infty} 2^{-k} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{2^{-k-1} \leq \Gamma \leq 2^{-k}\}} \right]$$

Key steps for the proof

Cut the sum according to the size of Γ_{ij} .

Set $F^t = M_n + M_n^t$ and $k_0 = \min\{k \geq 0; r_0 \leq 2^{-k-1}\}$.

$$\begin{aligned} \mathcal{I} &\leq \sum_{k=0}^{k_0} 2^{-k} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{\Gamma \geq r_0\}} \right] + \sum_{k=k_0+1}^{\infty} 2^{-k} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{2^{-k-1} \leq \Gamma \leq 2^{-k}\}} \right] \\ &\leq 2 \sum_{D \in \mathcal{C}(r_0)} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{I, I^t \in D\}} \right] + \sum_{k \geq 0} r_0 \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{2^{-k-1} \leq \Gamma \leq 2^{-k}\}} \right] \end{aligned}$$

Key steps for the proof

Cut the sum according to the size of Γ_{ij} .

Set $F^t = M_n + M_n^t$ and $k_0 = \min\{k \geq 0; r_0 \leq 2^{-k-1}\}$.

$$\begin{aligned}\mathcal{I} &\leq \sum_{k=0}^{k_0} 2^{-k} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{\Gamma \geq r_0\}} \right] + \sum_{k=k_0+1}^{\infty} 2^{-k} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{2^{-k-1} \leq \Gamma \leq 2^{-k}\}} \right] \\ &\leq 2 \sum_{D \in \mathcal{C}(r_0)} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{I, I^t \in D\}} \right] + \sum_{k \geq 0} r_0 \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{2^{-k-1} \leq \Gamma \leq 2^{-k}\}} \right] \\ &= 2 \sum_{D \in \mathcal{C}(r_0)} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{I, I^t \in D\}} \right] + r_0 \mathbb{E} \left[e^{\theta F^t/2} \right].\end{aligned}$$

Key steps for the proof

$$\mathbb{E} \left[e^{\theta F^t/2} \right] \leq \mathbb{E}[e^{\theta M_n}] \text{ Cauchy-Schwarz, Gaussian rotational invariance}$$

Key steps for the proof

$\mathbb{E} \left[e^{\theta F^t/2} \right] \leq \mathbb{E}[e^{\theta M_n}]$ Cauchy-Schwarz, Gaussian rotational invariance

Use of Hölder's inequality

$$\begin{aligned} \mathbb{E} \left[e^{\theta F^t/2} \mathbf{1}_{\{I, I^t \in D\}} \right] &= \mathbb{E} \left[e^{\theta M_n/2} \mathbf{1}_{\{I \in D\}} e^{\theta M_n^t/2} \mathbf{1}_{\{I^t \in D\}} \right] \\ &\leq \mathbb{E} \left[e^{\theta p M_n/2} \mathbf{1}_{\{I \in D\}} \right]^{1/p} \mathbb{E} \left[e^{\theta q M_n^t/2} \mathbf{1}_{\{I^t \in D\}} \right]^{1/q}. \end{aligned}$$

Key steps for the proof

By hypercontractivity,

$$\mathbb{E} \left[e^{\theta q M_n^t / 2} \mathbf{1}_{\{I^t \in D\}} \right]^{1/q} = \mathbb{E} [Q_t^q (e^{\theta M_n / 2} \mathbf{1}_{\{I \in D\}})]^{1/q}$$

Key steps for the proof

By hypercontractivity,

$$\mathbb{E} \left[e^{\theta q M_n^t / 2} \mathbf{1}_{\{I^t \in D\}} \right]^{1/q} = \mathbb{E} [Q_t^q (e^{\theta M_n / 2} \mathbf{1}_{\{I \in D\}})]^{1/q} \leq \mathbb{E} \left[e^{\theta p M_n / 2} \mathbf{1}_{\{I \in D\}} \right]^{1/p}$$

then

Key steps for the proof

By hypercontractivity,

$$\mathbb{E} \left[e^{\theta q M_n^t / 2} \mathbf{1}_{\{I^t \in D\}} \right]^{1/q} = \mathbb{E} [Q_t^q (e^{\theta M_n / 2} \mathbf{1}_{\{I \in D\}})]^{1/q} \leq \mathbb{E} \left[e^{\theta p M_n / 2} \mathbf{1}_{\{I \in D\}} \right]^{1/p}$$

then

$$\mathbb{E} \left[e^{\theta F^t / 2} \mathbf{1}_{\{I, I^t \in D\}} \right] \leq \mathbb{E} \left[e^{\theta p M_n / 2} \mathbf{1}_{\{I \in D\}} \right]^{2/p}$$

Key steps for the proof

By hypercontractivity,

$$\mathbb{E} \left[e^{\theta q M_n^t / 2} \mathbf{1}_{\{I^t \in D\}} \right]^{1/q} = \mathbb{E} \left[Q_t^q (e^{\theta M_n / 2} \mathbf{1}_{\{I \in D\}}) \right]^{1/q} \leq \mathbb{E} \left[e^{\theta p M_n / 2} \mathbf{1}_{\{I \in D\}} \right]^{1/p}$$

then

$$\mathbb{E} \left[e^{\theta F^t / 2} \mathbf{1}_{\{I, I^t \in D\}} \right] \leq \mathbb{E} \left[e^{\theta p M_n / 2} \mathbf{1}_{\{I \in D\}} \right]^{2/p}$$

by Hölder's inequality

$$\mathbb{E} \left[e^{\theta F^t / 2} \mathbf{1}_{\{I, I^t \in D\}} \right] \leq \mathbb{P}(I \in D)^{\frac{2-p}{p}} \mathbb{E} \left[e^{\theta M_n} \mathbf{1}_{\{I \in D\}} \right].$$

Finally with the property of the covering $\mathcal{C}(r_0)$.

$$\mathcal{I} \leq (r_0 + C\rho(r_0)^{\frac{2-p}{p}}) \mathbb{E}[e^{\theta M_n}].$$

Similar results

- ▶ Gaussian stationary processes on \mathbb{R}^d .

Similar results

- ▶ Gaussian stationary processes on \mathbb{R}^d .
- ▶ discrete Gaussian free field on \mathbb{Z}^d , $d \geq 3$

Similar results

- ▶ Gaussian stationary processes on \mathbb{R}^d .
- ▶ discrete Gaussian free field on \mathbb{Z}^d , $d \geq 3$
(infinite group of finite type+volume growth condition ok).

Similar results

- ▶ Gaussian stationary processes on \mathbb{R}^d .
- ▶ discrete Gaussian free field on \mathbb{Z}^d , $d \geq 3$
(infinite group of finite type+volume growth condition ok).
- ▶ Uniform measure on the sphere $\mathcal{S}^{n-1} \subset \mathbb{R}^n$.

Similar results

- ▶ Gaussian stationary processes on \mathbb{R}^d .
- ▶ discrete Gaussian free field on \mathbb{Z}^d , $d \geq 3$
(infinite group of finite type+volume growth condition ok).
- ▶ Uniform measure on the sphere $\mathcal{S}^{n-1} \subset \mathbb{R}^n$.
- ▶ Log-concave measures on \mathbb{R}^n with convexity assumptions.

Hypercontractivity relevant Gaussian processes \simeq
independent case (variance $\sim \frac{1}{\log n}$).

Hypercontractivity relevant Gaussian processes \simeq
independent case (variance $\sim \frac{1}{\log n}$).

discrete Gaussian free field on \mathbb{Z}^2 completely different behavior.

- ▶ $\text{Var}(M_n) = O(1)$ [Bramson-Ding-Zeitouni]

Hypercontractivity relevant Gaussian processes \simeq
independent case (variance $\sim \frac{1}{\log n}$).

discrete Gaussian free field on \mathbb{Z}^2 completely different behavior.

- ▶ $\text{Var}(M_n) = O(1)$ [Bramson-Ding-Zeitouni]
- ▶ convergence in distribution Gumbel randomly shifted [Bramson-Ding-Zeitouni '15].

Hypercontractivity relevant Gaussian processes \simeq
independent case (variance $\sim \frac{1}{\log n}$).

discrete Gaussian free field on \mathbb{Z}^2 completely different behavior.

- ▶ $\text{Var}(M_n) = O(1)$ [Bramson-Ding-Zeitouni]
- ▶ convergence in distribution Gumbel randomly shifted [Bramson-Ding-Zeitouni '15].

Hypercontractivity alone doesn't work .

Thank you for your attention.